ERF 26th Annual Conference

The Dynamic Interaction of Exchange Rates and International Trade Flows in the MENA Region: GARCH Analysis

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Agenda







1. Introduction

The role of exchange rates and their impacts on macroeconomic performance are subjects of debate in the economics literature and core areas of interest to economists, finance experts, and policy makers.

Most economists agree that exchange rate stability enhances foreign investment, international trade, and economic growth.

Research work on exchange rate volatility and international trade in MENA region is scanty. Despite the existence of an abundance of literature on international trade inflows, exchange rate, and growth, there is still a paucity of empirical studies on the dynamic interaction between these aforementioned variables.



2. Selected Literature Review on Exchange Rate and International Trade

- Exports are more sensitive to changes in REER levels than their volatility, and the effect is more pronounced in the agriculture sector exports. Huchet-Bourdon and Korinek (2011)
- High external debt could offset the effect on trade of exchange rate movements. Kearns and Patel (2016)
- Developing countries should be cautious about relaxation in the exchange rate and embrace of the floating exchange rate, which can be dangerous for international trading performance without a highly developed financial system. Héricourt and Poncet (2015)

Shocks causing changes in the exchange rate can lead to changes in other macroeconomic variables, which could offset the impact of exchange rate on trade. Bacchetta and van Wincoop (2000)

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- Baek studied exports and imports of 71 products between the US and the Republic of Korea and concluded that exchange rate levels affect exports and imports between the Republic of Korea and the US. Baek (2012)
- The depreciation of the US dollar had led to exchange rate appreciation in developed Asia, and that this in turn might disrupt the complementary trade relationship between developed Asia and developing economies in the region. Thorbecke (2006)
- Currency depreciation led to lower export prices paid by foreigners and higher import prices, and these price changes in turn led to a rise in exports and a fall in imports. An increase in exports is higher when the exporting economy has a weaker financial system, especially in cases of banking crises. IMF (2015)

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3. Recent Developments in the MENA Region

GDP Growth

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Sources: Haver Analytics, IMF, World Bank

Tourism Growth



Sources: Haver Analytics, IMF, World Bank

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Good Exports Growth: Oil Importers



Sources: Haver Analytics, IMF, World Bank

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World Bank Human Capital Index



Sources: Haver Analytics, IMF, World Bank

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Inflation

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Sources: Haver Analytics, IMF, World Bank

GCC Interest Rate

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Sources: Haver Analytics, International Country Risk Guide, IMF, World Bank.

Political Stability and Business Climate



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Sources: Haver Analytics, International Country Risk Guide, IMF, World Bank.

Export Market Penetration Index



Sources: Haver Analytics, International Country Risk Guide, IMF, World Bank.

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Euro Area Trade Exposure

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Sources: Haver Analytics, International Country Risk Guide, IMF, World Bank.



4. Methodology

Measures for Exchange Rate Volatility

In this study, exchange rate volatility is measured by applying the symmetric GARCH model and asymmetric threshold-GARCH (TGARCH) model, The simplest forms of the GARCH model and TGARCH model are GARCH (1,1) and TGARCH (1,1), which can be shown as follows:

$$R_{t} = \beta_{0} + \beta_{1}R_{t-1}$$
(1)

$$\sigma_{t}^{2} = \omega + \alpha \mu_{n-1}^{2} + \beta \sigma_{n-1}^{2}$$
(2)

$$\sigma_{t}^{2} = \omega + \alpha \mu_{n-1}^{2} + \beta \sigma_{n-1}^{2} d_{n-1} + \gamma \sigma_{n-1}^{2}$$
(3)

where ω is constant. $\boldsymbol{\alpha}$, β and γ are coefficients. μ_{n-1}^2 is the mean square of the previous time period. σ_{n-1}^2 is the variance of the previous time period. *dn* takes the value of 1 for $\mu_t < 0$, and 0 otherwise.

Model Specification

To investigate the effect of exchange rate volatility on international trade, equation (4) below is estimated, incorporating pregenerated exchange rate volatility proxies.

 $Trade_{it} = \beta_0 + \beta_1 EXVOL_{it} + \beta_2 ECOFR_{it} + \beta_3 INFL_{it} + \beta_4 INTR_{it} + \beta_5 IND_{it} + \beta_6 RES_{it} + \epsilon_{it}$ (4)

In Equations (4), for country i and time period t, $TRADE_{it}$ represents international, $EXVOL_{it}$ is exchange rate volatility, $ECOFR_{it}$ is economic freedom, $INFL_{it}$ is inflation rate, $INTR_{it}$ is interest rate, IND_{it} is industrial growth, RES_{it} is reserve and ϵ_{it} denotes to error term.

Methodology Map







5. Sampling and Data Collection

Sample

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This study uses a sample of **15 MENA countries**: Algeria, Bahrain, Djibouti, Egypt, Iraq, Iran, Jordan, Kuwait, Lebanon, Morocco, Oman, Qatar, Saudi Arabia, Tunisia and the United Arab Emirates. *Libya, Syria, West Bank, and Gaza and Yemen are excluded from the study due to data unavailability.*

Annual time series data (1990-2018) were sourced from the Statistical Bulletin of Central Banks, World Development Indicators (WDI) from the World Bank data set and International Financial Statistics (IFS) developed by the IMF.

Description of Variables and Sources of Data

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Variables	Description	Data Source		
TRADE	International trade measured by total	World Bank Development Indicators		
	trade (% of GDP)			
	The volatility of real effective	Bruegel Datasets: Real effective		
EXVOL	exchange rate (%), measured by	exchange rates for 178		
	GARCH (1,1) & TGARCH (1,1)	countries: A new database		
ECOFR	Economic freedom index	U.S. Heritage Foundation,		
INFL	Inflation rate measured by consumer	IMF IFS Statistics,		
	price index			
INTR	Interest rate measured by the real	World Bank Development Indicators		
	interest rate			
IND	Industrial growth measured by	World Bank Development Indicators		
<u></u>	value-added growth			
RES	Official reserve	World Bank Development Indicators		









6. Results7. Interpretation of the findings

Unit Root Test

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	Level				First Difference			
Variable	Intercept Interce		Intercep	pt and Trend In		ercept	Intercept and Trend	
	Prob.*	T-Statistics	Prob.*	T-Statistics	Prob.*	T-Statistics	Prob.*	T-Statistics
TRADE	0.4713	-2.971853	0.9122	-3.580623	0.0005*	-2.976263	0.0024*	-3.587527
EXVOL	0.6329	-1.261421	0.1022	-3.242082	0.0009*	-3.745469	0.0448*	-3.641270
INFL	0.5207	-2.971853	0.6650	-3.580623	0.0002*	-5.392448	0.0009*	-5.403225
INTR	0.0137*	-2.971853	0.0228*	-3.953571	0.0000*	-7.190693	0.0000*	-7.274761
IND	0.0107	-3.659743	0.0352*	-3.750388	0.0000*	-8.384662	0.0000*	-8.239793
RES	0.2322	-2.138279	0.0352*	-3.801457	0.04392*	-2.245913	0.04296*	-2.425841
ECOFR	0.5562	-1.422841	0.6878	-1.7776691	0.0420*	-2.871084	0.0184*	-2.879148

*Result is significant at 5% level

Source: Authors' estimation (statistical work is performed using E-views software version 9).

Johansen Cointegration Test

Sample (adjusted): 1992 2018 Included observations: 27 after adjustments Trend assumption: Linear deterministic trend Series: TRADE ERVOL ECOFR INFL IND INTR RES Lags interval (in first differences): 1 to 1 Unrestricted Cointegration Rank Test (Trace)

Hypothesized No. of CE(s)	Eigenvalue	Trace Statistic	0.05 Critical Value	Prob.**
None *	0.935904	206.9023	125.6154	0.0000
At most 1 *	0.832029	132.7234	95.75366	0.0000
At most 2 *	0.707396	84.55634	69.81889	0.0021
At most 3 *	0.594869	51.37510	47.85613	0.0225
At most 4	0.416446	26.97938	29.79707	0.1021
At most 5	0.246283	12.43668	15.49471	0.1372
At most 6 *	0.162956	4.802731	3.841466	0.0284

Trace test indicates 4 cointegrating eqn(s) at the 0.05 level

* denotes rejection of the hypothesis at the 0.05 level

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Source: Author's estimation (statistical work is performed using E-views software version 9)

Vector Error Correction Model (VECM)

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Vector Error Correction Esti	mates		R-squared	0.693092	0.816579
Sample (adjusted): 1993 201	8		Adj. R-squared	0.488487	0.694299
Included observations: 26 af	ter adjustments		Sum sq. resids	142.4833	415.5152
Standard errors in () & t-sta	tistics in []		S.E. equation	3.082027	5.263175
Cointegrating Eq:	CointEq1		F-statistic	3.387464	6.677922
			Log likelihood	-59.00707	-72.92090
TRADE(-1)	1.000000		Akaike AIC	5.385159	6.455453
ERVOL(-1)	-0.620375		Schwarz SC	5.917431	6.987725
× ź	(0.77388)		Mean dependent	0.287954	-1.601298
	[-0.80164]		S.D. dependent	4.309314	9.519174
С	9.422432				010 (000
			Determinant resid covaria	nce	213.4220
Error Correction:	D(TRADE)	D(ERVOL)	Determinant resid covaria	nce	71.03541
((fallming) States			Log likelihood		-129.2061
			Akaike information criteri	on	11.78509
- Office and a state of the sta			Schwarz criterion		12.94641

Source: Authors' estimation (statistical work is performed using E-views software version 9)

Impulse Response Function (IRF)

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Source: Author's estimation (statistical work is performed using E-views software version 9)

Heteroskedasticity Test: ARCH for Exchange Rate Volatility

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F-statistic	63.33764	Prob. F(1,26)	0.0000
Obs*R-squared	19.85114	Prob. Chi-Square(1)	0.0000

Source: Authors' estimation (statistical work is performed using E-views software version 9).

Estimation of Parameters of GARCH/TGARCH Model

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	Variable	Coefficient	Std. Error	z-Statistic	Prob.			
	ERVOL	-0.291627	0.035393	-8.239640	0.0000**			
Dependent Variable: IRADE	ECOFR	0.020606	0.006904	2.984613	0.0028**			
Method: ML ARCH - Normal distribution (BFGS / Marguardt steps)	INFL	0.009709	0.001216	7.985518	0.0000**			
Sample: 1990 2018	IND	3.121113	1.285931	2.427124	0.0152**			
Included observations: 20	INTR	-0.058756	0.098273	-0.597884	0.5499			
Included observations: 29	RES	5.52E-10	1.20E-10	4.618867	0.0000**			
Convergence not achieved after 500 iterations								
Coefficient covariance computed using outer product of gradients		Variance Equation						
Presample variance: backcast (parameter $= 0.7$)	С	36.69950	31.36451	1.170096	0.2420			
$GARCH = C(6) + C(7)*RESID(-1)^{2} + C(8)*GARCH(-1)$	RESID(-1)^2	-0.172162	0.175904	-0.978730	0.3277			
	GARCH(-1)	0.622406	0.446011	1.395494	0.1629			
	R-squared	-0.139929	Mean dependent	var	59.66699			
	Adjusted R-squared	-0.329917	 7 S.D. dependent var 9 Akaike info criterion 7 Schwarz criterion 		8.801952			
and and an and a second and a second and a second a secon	S.E. of regression	10.15059			7.561375			
mittillion from the second	Sum squared resid	2472.827			7.938560			
	Log likelihood	-101.6399	Hannan-Quinn ci	riter.	7.679505			
and the state of t	Durbin-Watson stat	1.417027						
Significance	e Levels *** p < 0 <u>.01</u> , **	*p<0.05,*	p < 0.1.					

Source: Authors' estimation (statistical work is performed using E-views software version 9).

Estimation of Parameters of EGARCH Model

Dependent Variable: TRADE	Variable	Coefficient	Std. Error	z-Statistic	Prob.
Method: ML ARCH - Normal distribution (BFGS / Marquardt steps) Sample: 1990 2018 Included observations: 29 Convergence not achieved after 500 iterations Coefficient covariance computed using outer product of gradients Presample variance: backcast (parameter = 0.7)	ERVOL ECOFR INFL IND INTR RES	-0.292874 0.019939 4.017838 3.311586 -0.078894 5.35E-10	0.016020 0.004513 1.419821 0.623985 0.030472 8.67E-11	-18.28207 4.418292 2.829820 5.307152 -2.589043 6.170208	0.0000** 0.0000** 0.0047** 0.0000** 0.0096** 0.0000**
LOG(GARCH) = C(6) + C(7)*ABS(RESID(-1)/@SQR1(GARCH(-1))) + C(8) *RESID(-1)/@SQRT(GARCH(-1)) + C(9)*LOG(GARCH(-1))		Variance	Equation		
	C(6)	4.553736	1.158969	3.929127	0.0001
	C(7)	-1.916115	0.803926	-2.383446	0.0172
	C(8)	0.411485	0.395884	1.039410	0.2986
	C(9)	0.256958	0.179998	1.427560	0.1534
	R-squared	-0.178086	Mean dependent	var	59.66699
and the second sec	Adjusted R-squared	-0.374434	S.D. dependent v	ar	8.801952
St Halling Reasons Co.	S.E. of regression	10.31908	Akaike info crite	rion	7.371313
	Sum squared resid	2555.601	Schwarz criterion		7.795646
	Log likelihood	-97.88403	Hannan-Quinn cr	iter.	7.504208
Botting	Durbin-Watson stat	1.402172			
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Significance Levels *** p < 0.01, ** p < 0.05, * p < 0.1. **Source:** Author's estimation (statistical work is performed using E-views software version 9).

Conclusion





The coefficients of exchange rate volatility measured with TGARCH/EGARCH showed significant negative effects in both cases. This indicates that high exchange rate volatility can adversely affect international trade flows in sample countries.

Economic freedom - as a proxy of financial market development and economic openness - has a significant and positive impact on international trade .



Inflation has a significant positive effect on international trade as an explanatory variable, as does exchange rate volatility.



The statistical significance of both industrial growth and reserves boosted international trade. This indicates that effective import substitution and industrialization would significantly reduce pressure on the external sector and will increase economic activities and hence, spur economic growth.

The findings revealed that exchange rate volatility depressed trade flows in the long run.

Policy Implication

This study suggests that economic policies of MENA countries should be designed to minimize exchange rate volatility. The minimum level of exchange rate volatility could be beneficial to foster international trade inflows, which might increase the overall economies of these countries. Future studies should analyze the relationship between trade with exchange rate volatility along with other macroeconomic, social, and political factors in developed and developing countries, which might further guide policy directions.



Summery

Panel annual time series **data** (1990-2018) **MENA** **CVS** are: economic freedom; inflation rate; interest rate; reserves; and industrial development Economic policies of MENA countries should be redesigned to minimize exchange rate volatility and therefore, foster trade

This study examined the dynamics between real exchange rate volatility and the level of international trade



GARCH (1,1)

Model is employed to measure ER volatility **Results** revealed that, in the long-run, exchange rate volatility negatively affects international trade significantly in the sample countries



